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1 **Efficacy and safety of anticoagulants in patients with atrial fibrillation and history of falls or risk of**
2 **falls. A systematic review and multilevel meta-analysis**

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20 **Running head:** Multilevel meta-analysis of anticoagulants in patients with atrial fibrillation and falls.

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35 Not applicable.

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41 Template data collection form, data extracted, analytical code and any other study documents will be
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43 **Code availability**

44 Analytical code will be available from the corresponding author on reasonable request.

45 **Authors' contributions**

46 Conceptualization: GYHL, PL, GC, SD, RH and TG; methodology: GYHL, PL, GC, SD, RH and TG; formal
47 analysis: GC and TG; investigation: GYHL, PL, GC, SD, RH and TG; data curation: GC and TG; writing

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54 **Abstract**

55 **Introduction:** Atrial fibrillation (AF) is a major cause of stroke. Anticoagulants substantially reduce
56 risk of stroke but are also associated with an increased risk of bleeding. Because of that many
57 patients do not receive anticoagulants; particularly patients at risk of falls. This systematic review and
58 meta-analysis aims to compare anticoagulant treatment options for the management of atrial
59 fibrillation patients at risk of falls or with a history of falls.

60 **Methods:** We conducted a PRISMA systematic review (until March 2022), including studies
61 evaluating safety and efficacy of different anticoagulants (Vitamin K antagonist [VKA] versus non-
62 vitamin K antagonist oral anticoagulant [NOAC]). Outcomes were ischemic stroke, major bleeding,
63 intracranial hemorrhage, hemorrhagic stroke, myocardial infarction, gastro-intestinal bleeding,
64 cardiovascular and all-cause mortality. A multilevel meta-analysis was conducted adjusting for
65 clustering effects within studies examining more than one effect size.

66 **Results:** 919 articles were identified, 848 after removing duplicates. 155 were screened for full text
67 and 10 articles were retained for final quantitative synthesis. Risk of bias was moderate to serious for
68 the included studies. In meta-analysis, NOACs were associated with superior effectiveness compared
69 to VKA for ischemic stroke/systemic embolism (HR 0.82, 95%CI [0.69–0.98]; $p < 0,05$) and safety
70 (hazard ratio (HR) 0.53, 95% confidence interval (CI) [0.40–0.71]; $p < 0,05$) for intracranial
71 hemorrhage. There were no differences in other outcomes.

72 **Conclusion:** NOACs were associated with less intracranial hemorrhages and ischemic
73 strokes/systemic embolisms than VKAs in AF patients at risk of falls. These findings suggesting
74 preferred use of NOACs over VKAs would have clinical implications for physicians, patients and policy
75 makers.

76 **Key points:**

77 • AF patients at risk or with history of falls often do not receive anticoagulants. Anticoagulation
78 treatments for AF patients at risk or with history of falls is an under-researched area and
79 clinical guidelines are missing.

80 • This systematic review and multilevel meta-analysis evaluated safety and efficacy of NOACs
81 compared to VKA in patients with atrial fibrillation and at risk of falls or with history of falls.

82 • NOACs were associated with less intracranial hemorrhages and ischemic strokes/systemic
83 embolisms than VKAs in AF patients at risk of falls. These findings suggesting preferred use of
84 NOACs over VKAs would have clinical implications for physicians, patients and policy makers.

85

86

87 **1. Introduction**

88 Atrial fibrillation (AF) is the most common sustained cardiac arrhythmia [1] and is a major cause of
89 stroke, heart failure, and death [2], as well as healthcare costs [3]. Stroke is the second most
90 common cause of death in the general population and it is a major cause of disability [4,5]. AF
91 patients have a yearly risk of stroke of 5%, and this risk is increased in the presence of certain risk
92 factors, including left ventricular dysfunction, hypertension, a history of stroke, and increasing age
93 [6].

94 Treatment with oral anticoagulants (OACs) substantially reduces risk of stroke but is also associated
95 with an increased risk of bleeding and especially intracranial hemorrhage which is the most feared
96 complication [7,8]. Because of that, many patients do not receive anticoagulants, and particularly
97 patients at risk of falls or with history of falls [9,10]. In eligible patients with elevated stroke risk
98 overall, the median rate of non-treatment is 23.3% (from 7.9% to 51.1%) [11]. In patients at risk of
99 falls, this rate was estimated at 50% in the era pre-NOAC [12]. AF patients at high risk of falls and on
100 oral anticoagulation do not have a significantly increased risk of major bleeding, suggesting that
101 being at risk of falls should not prevent OAC prescribing [13,14]. Additionally, the HAS-BLED risk
102 stratification tool for bleeding assessment in anticoagulated patients with atrial fibrillation does not
103 consider falls (risk or history) as an independent predictor of major bleeding [15,16].

104 The non-vitamin K antagonist oral anticoagulants (NOACs) were shown in a number of systematic
105 reviews and meta-analyses to reduce the risk for intracranial hemorrhage by approximately 50%
106 compared to vitamin K antagonists (VKAs) in the general AF population at risk of stroke [17,18], and
107 are therefore the preferred option in guidelines [19,20]. NOACs might be the most appropriate
108 anticoagulant in patients with an increased risk of falls and help to alleviate fears of bleeding
109 complications.

110 To our knowledge there is limited evidence and there are no recommendations and guidelines for
111 the use of NOACs specifically for the patients at risk of falls or with history of falls. To date, efforts

112 have been focusing on elderly patients only. In a recent meta-analysis in older AF patients, NOACs
113 were associated with superior efficacy in preventing stroke/systemic embolism (hazard ratio [HR]
114 0.83, 95% CI: 0.74-0.94), superior safety for intracranial bleeding (HR 0.58, 95% CI: 0.50-0.67) and
115 non-inferiority safety for major bleeding (HR 0.93, 95% CI: 0.86-1.01) and gastrointestinal bleeding
116 (HR 1.17, 95% CI: 0.99-1.38) compared to VKAs [21]. Whether NOACs are the most appropriate
117 anticoagulant treatment option for AF patients at risk of falls or with history of falls remain uncertain.

118 In the first contemporary study of its kind, the Liverpool AF-Falls project aims to determine the safety
119 and efficacy of NOACs compared to vitamin K antagonists (VKAs) for the management of AF patients
120 at risk of falls or with a history of falls. Results from the project could provide clinicians and policy
121 makers with information on which to make evidence-based recommendations.

122

123 **2. Methods**

124 The protocol has been registered in the International Prospective Register of Systematic Reviews
125 (PROSPERO) database (CRD42020201086) [22]. The methodology used for this systematic review
126 follows the recommendations of the Cochrane Handbook for Systematic Reviews of Interventions
127 [23]. This review is reported according to PRISMA 2020 and the checklist is available as an additional
128 file (electronic supplementary material [ESM] Table S1) [24].

129

130 **2.1. Eligibility criteria**

131 Randomized controlled trials (RCT) (including *post hoc* and ancillary analysis), quasi-randomized
132 studies and observational (prospective, retrospective, case control and cohort studies) studies were
133 included. Animal studies, editorials, letters, case reports, reviews, case series, eminence-based
134 opinions and conference abstracts were excluded. Systematic reviews of interventions were
135 excluded but included studies from relevant systematic reviews were assessed for inclusion.

136 We included studies of adults (age 18 or older) patients with any forms of nonvalvular AF
137 (paroxysmal, persistent or permanent) with history of falls or that are at risk of falls comparing
138 NOACs to VKAs. Patients were defined at risk of falls if they had one of these criteria based on a
139 revised list from Steffel *et al.* [25]: prior history of falls; lower extremity weakness; poor balance;
140 cognitive impairment; vision and/or hearing impairment; orthostatic hypotension; use of
141 psychotropic or antihistaminic, or anticholinergic, or antihypertensive drugs; severe arthritis;
142 dizziness; frailty; polypharmacy defined as a minimum of six pharmaceutical treatments and
143 multimorbidity defined as a minimum of four comorbidities.

144 Studies including patients receiving ablation, cardioversion, or left-atrial appendage closure were
145 excluded.

146 The primary efficacy outcome was the composite of ischemic stroke and/or systemic embolism (an
147 acute vascular occlusion of an extremity or organ). The primary safety outcome was major bleeding
148 (defined based on International Society on Thrombosis & Haemostasis for major bleeding in non-
149 surgical patients) [26]. Secondary outcomes included: intracranial hemorrhage (including all
150 intracerebral, subdural, epidural, subarachnoid hemorrhage and hemorrhagic stroke);
151 gastrointestinal bleeding; clinically relevant non-major bleeding (defined based on International
152 Society on Thrombosis & Haemostasis for major bleeding in non-surgical patients) [27]; myocardial
153 infarction; ischemic stroke; systemic embolism; hemorrhagic stroke; cardiovascular mortality and all-
154 cause mortality.

155 **2.2. Search methods for identification of studies**

156 The following bibliographic databases were searched: Cochrane Central Register of Controlled Trials
157 (CENTRAL), CINAHL, Embase (via OVID); MEDLINE (via OVID), Scopus and Web of Science. We also
158 searched the following trials register: the US National Institutes of Health Register
159 (www.clinicaltrials.gov). Finally, we double-checked the reference lists of all the relevant studies and
160 reviewed the articles to identify additional relevant studies. English-language articles published from

161 inception to March 2022 were identified. Regular alerts were also established to identify subsequent
162 publications.

163 The search strategy for bibliographic databases was developed from the research question and
164 implemented by a health sciences librarian with expertise in searches for systematic reviews. A
165 combination of terms of medical subject headings (MeSH) and keywords was used in the search
166 strategy for MEDLINE (ESM Table S2). For Embase, similar terms and search limits were used. MeSH
167 terms were replaced with Emtree indexing terms and/or keywords, as appropriate. The search
168 strategies for MEDLINE and Embase was adapted for use in Scopus, Web of Science and the other
169 bibliographic databases. The search results were entered into the EndNote X8 reference
170 management software for screening, once duplicate records were removed using EndNote X8.

171 **2.3. Data collection and analysis**

172 2.3.1 Selection of studies

173 Two independent reviewers (TG and GC) performed study selection. During stage 1, titles and
174 abstracts were screened to identify potentially relevant studies applying the inclusion and exclusion
175 criteria. At stage 2, full-text review established the final set of included studies, with discrepancies
176 resolved by a third reviewer (GL). The reason for exclusion were noted for all articles rejected at
177 stage 2. Study authors were contacted in cases further information was needed to make a screening
178 decision. A PRISMA flow diagram was developed to record the study selection process [24].

179 Data was extracted from each eligible study using a custom data extraction template by one reviewer
180 (TG) and cross-checked with the source article by a second reviewer (GC). Discrepancies and
181 differences in interpretation were resolved through discussion, and if necessary, by consultation with
182 a third reviewer (PL or GL). Where insufficient data were presented, we requested additional
183 information from the study authors. The following were collected from each study: study
184 characteristics (publication year, authors, title, study objectives and study outcomes), study
185 population (such as age, gender, and diagnostic criteria), study design, intervention and control

186 details, and outcomes (hazard ratios, standard error or 95% confidence intervals). For observational
187 studies, adjusted results were preferred over non adjusted, when available.

188 2.3.2 Assessment of risk of bias in included studies

189 In this systematic review, risk of bias in observational studies was appraised with the Risk Of Bias In
190 Non-randomized Studies - of Interventions I tool (ROBINS-I tool) [28]. Using this tool, studies are
191 scored as low, moderate, serious or critical risk of bias. Confounding domains included
192 demographics, comorbidities, bleeding risk, stroke risk and concomitant treatments. Co-
193 interventions included anti-platelet agents. The effect of interest was the effect of assignment. Risk
194 of bias was independently evaluated by two reviewers (TG and GC) and we resolved any
195 disagreements with a third reviewer (GL).

196 **2.4. Statistical Analysis**

197 Data synthesis was conducted based on the sufficient clinical homogeneity regarding participant
198 characteristics, types of intervention and outcomes, and comparability between methods and ability
199 to aggregate data. Statistical heterogeneity as consequence of clinical and/or methodological
200 diversity was evaluated both by visual inspection of the forest plots and a formal statistical test, using
201 Cochran Q test and I^2 statistic [29]. If heterogeneity was low or minor, a fixed effect model was used
202 to pool the data; if heterogeneity was moderate-to-substantial a random-effects model was used
203 instead.[29] For the fixed effect model, the generic inverse variance method was used. For the
204 random-effects model, data was pooled across studies using the DerSimonian and Laird model [30].
205 For outcomes that included studies with multiple effect sizes (e.g., when a study provided separate
206 effect sizes for different NOACs, or different subgroups of patients being at risk of falls) a multilevel
207 random effects meta-analysis was conducted which takes into account the hierarchical structure of
208 the dataset [31,32]. We assumed that effect sizes within studies were correlated with a correlation
209 coefficient $\rho=0.5$ to calculate the variance-covariance matrix (sensitivity analyses were conducted
210 using $\rho=0.3$ and $\rho=0.7$) [32,33]. The restricted maximum likelihood method was used to estimate

211 model overall effect. Confidence intervals of the model coefficients were calculated with robust
212 variance estimation [34].

213 Results of meta-analysis were presented as pooled HRs with 95% CIs. We assessed the publication
214 bias using Funnel plot and Egger's regression test for outcomes reported in at least 5 studies. A
215 modified version of the Egger's regression test was used with robust variance estimation for handling
216 dependency for outcomes that included studies with multiple effect sizes [35]. Statistical analyses
217 were performed in R using RStudio version 4.0.0 (meta and metafor packages) [36].

218 **2.5. Quality of evidence**

219 Two reviewers (TG and GC) assessed the quality of evidence with the Grading of Recommendations
220 Assessment, Development and Evaluation (GRADE) system which considers study design, risk of bias,
221 inconsistency of results, indirectness, imprecision and other factors [37]. GRADE Guidelines 18 was
222 followed as we assessed non-randomized studies with the ROBINS-I tool [38]. Disagreements were
223 resolved by a third review author (PL or GL). Assessment results were described in summary of
224 findings (SoF) tables; GRADEpro GDT was used to create SoF tables [39].

225 **3. Results**

226 **3.1. Study selection and characteristics**

227 During the search process, 919 abstracts were identified. Following the removal of duplicates, 693
228 abstracts were excluded at stage 1, 155 full-text articles were assessed further for eligibility, and 10
229 met eligibility criteria for inclusion in this review [25,40–48]. Reasons for exclusion included study
230 design (22 articles), comparator (15 articles), outcome (six articles), population (80 articles), and
231 research question (22 articles). The PRISMA flow diagram is presented in Fig. 1. All studies were non-
232 randomized. Five articles were retrospective cohorts [41,44,46–48], and the others were subgroup
233 analyses of randomized clinical trials (one pre-specified subgroup analysis of ENGAGE-AF-TIMI trial
234 [25,49], three post-hoc analyses of the ARISTOTLE trial [40,42,43,50] and one post-hoc analysis of the
235 ROCKET AF trial [45,51]). Lip *et al.* 2020 [46], Hohmann *et al.* 2019 [48] and Martinez *et al.* 2018 [47]
236 contributed three effect sizes each as they investigated either different NOACs compared to VKA, or

237 different subgroups of AF patients being at risk of falls. The subgroup analyses [40,42,43] of the
238 ARISTOTLE trial also contributed three effect sizes as they analyzed distinctly different subgroups of
239 AF patients being at risk of falls. Sample sizes ranged between 617 and 79,796 AF patients at
240 moderate or high thromboembolic risk and with history of falls [40] or at risk of falls [25,41–48].
241 Most of the patients were elderly, mean age ranging from 71 to 83. The median follow-up period for
242 studies ranged from 0.5 to 2.8 years. The main characteristics of the studies included are shown in
243 Table 1.

244 **3.2. Risk of bias of included articles**

245 Based upon the ROBINS-I tool for non-randomized studies, the overall risk of bias ranged from
246 moderate to serious according to the included articles (ESM Fig. S1). All studies conducted
247 appropriate statistical methods to adjust for the confounders (domain 1), not always integrating all
248 the confounding domains (3/10). They used either propensity score weighing (2/10) or matching
249 (1/10) methods or multivariate Cox regression model (7/10) to reduce confounders and account for
250 covariates that may impact the outcomes. Regarding selection bias and bias due to missing data
251 (domains 2 and 5), the included studies were at low risk. In relation to bias in classification of the
252 intervention (domain 3), bias due to deviation from the intended intervention (domain 4), bias in
253 measurement of outcome and reporting bias (domains 6 and 7), included studies were at low or
254 moderate risk of bias for these domains except Martinez *et al.* 2018 which was at serious risk of bias
255 due to deviation from the intended intervention as they did not censor follow-up time for treatment
256 switching [47].

257 **3.3. Outcome assessment**

258 In the prophylaxis of stroke or systemic embolism (15 effect sizes), NOACS were superior to VKAs
259 (hazard ratio (HR) 0.82, 95% confidence interval (CI) [0.69–0.98]; $p < 0.05$; $I^2 = 67.7\%$), Fig. 2. Of the 10
260 articles included in the meta-analysis, seven evaluated the hazard for intracranial hemorrhage (15
261 effect sizes), which was lower with NOACs compared to VKA (HR 0.53, 95%CI [0.40–0.71]; $p < 0.05$;
262 $I^2 = 46\%$), Fig. 3. In reducing the risk of major bleeding (11 effect sizes), NOACs were not different from

263 VKAs (HR 0.88, 95%CI [0.74–1.04]; p=0.09) (ESM Fig. S2). There were no differences between NOACs
264 and VKA regarding risks in ischemic stroke (HR 0.87, 95%CI [0.60-1.28], p=0.23; eight effect sizes),
265 hemorrhagic stroke (HR 0.51, 95%CI [0.24–1.10]; p=0.10; nine effect sizes), gastro-intestinal bleeding
266 (HR 1.04, 95%CI [0.89-1.23], p=0.44; 12 effect sizes), myocardial infarction (HR 0.76, 95%CI [0.47-
267 1.24], p=0.27; fixed effect model, reported in two studies), cardiovascular mortality (HR 1.04, 95%CI
268 [0.61–1.75]; p=0.89; random effect model, reported in two studies) and all-cause mortality (HR 1.23,
269 95%CI [0.35–4.29]; p=0.55; five effect sizes) (ESM Figures S3-S8). Sensitivity analyses results were
270 aligned with the main results, regardless of the outcomes, for a correlation coefficient $\rho=0.3$ or $\rho=0.7$
271 (ESM Table S4).

272 **3.4. Certainty of evidence**

273 A detailed quality assessment of study outcomes for the comparisons of NOACs to VKAs, where two
274 or more studies were available, is given in the summary of findings table (Table 2) and the GRADE
275 evidence profile (ESM Table S3). In all comparisons the quality of the evidence was low or very low
276 grade, according to Working Group GRADE of evidence. The quality of evidence was downgraded
277 according to the different outcomes because of some concerns in regard to the risk of bias due to
278 confounding and deviation from the intended intervention, inconsistency, imprecision and
279 publication bias. There was evidence of publication bias for the intracranial hemorrhage outcome
280 (Funnel plot asymmetry [ESM Fig. S9] and Egger's regression test, $p<0.01$), but not for ischemic
281 stroke/systemic embolism (ESM Fig. S10 and Egger's regression test, $p=0.06$), major bleeding (ESM
282 Fig. S11 and Egger's regression test, $p=0.26$) and gastro-intestinal bleeding (ESM Fig. S12 and Egger's
283 regression test, $p=0.68$). Publication bias was not investigated for the other outcomes due the limited
284 number of studies.

285 **4. Discussion**

286 This systematic review and meta-analysis of 10 studies is the first to compare NOACs to VKAs as
287 anticoagulation strategies for patients with non-valvular atrial fibrillation and at risk of falls or with
288 history of falls. The main findings from the pooled analyses were as follows: (1) there was a 18%

289 reduction in the risk of stroke or systemic embolism with NOACs compared to VKAs and a 47%
290 reduction in the risk of intracranial hemorrhage. (2) The risk of major bleeding events is not different
291 between groups, as were the risks for ischemic stroke, hemorrhagic stroke, gastro-intestinal
292 bleeding, myocardial infarction, cardiovascular and all-cause mortality. While we found a 18%
293 reduction in the risk of stroke or systemic embolism, there was no difference in ischemic stroke
294 alone. However, the effect sizes were similar. This could be explained due to sample size as ischemic
295 stroke rate alone is about 20-40% lower than ischemic stroke or systemic embolism rate.

296 Given the relatively modest improvement of NOACs in preventing thromboembolic events such as
297 stroke or systemic embolism compared to VKA, the safety of each treatment is of paramount
298 importance and must be rigorously considered to decide the most appropriate antithrombotic
299 management. We found in our meta-analysis a 47% reduction in the risk of intracranial hemorrhage
300 with NOACs as compared to VKAs. The shorter half-life of NOACs and the more targeted mechanism
301 of anticoagulation (direct thrombin or factor Xa inhibition) have been implicated in the reduction of
302 intracranial hemorrhage with these agents as compared to VKAs [49]. The 2014 AHA/ACC/HRS
303 Guideline for the Management of Patients With Atrial Fibrillation does not make specific
304 recommendations for use of anticoagulation in AF patients at risk of falls or with history of falls [52]
305 nor the 2019 update [53] with NOACs being preferred over VKAs overall. The 2020 European
306 guidelines suggests that the increased risk of bleeding in patients at risk of falls does not outweigh
307 the benefits of anticoagulants and suggests that NOACs have a better risk-benefit profile over VKAs
308 based on evidence from studies on elderly AF patients [20]. The present meta-analysis adds to the
309 limited body of evidence in AF patients at risk of falls or with history of falls suggesting that NOACs
310 may be the optimal strategy for antithrombotic management, given the improved efficacy in
311 preventing thromboembolic events and the improved safety profile as compared to VKAs. This study
312 confirms that NOACs have a better risk-benefit profile than VKAs, as it was shown in the broader AF
313 population or in elderly AF patients.

314 In the broader AF patients, large RCTs have shown that NOACs are at least as effective as VKAs for
315 preventing stroke and systemic embolization and are also associated with significantly less
316 intracranial hemorrhage and major bleeding events [49–51,54]. Meta-analyses of RCTs found NOACs
317 to be superior to warfarin for the prevention of stroke and systemic embolism [55–58].

318 However, there is less evidence for the use of NOACs in AF patients at risk of falls or with history of
319 falls. Related subgroup of AF patients made of older individuals, particularly those aged 75 years and
320 over, have been studied substantially [59–62]. Meta-analyses including RCTs and observational
321 studies showed that NOACs (i) were as effective as VKAs in reducing stroke and systemic embolism,
322 (ii) were not significantly different for major bleeding and (iii) significantly reduced risk of intracranial
323 hemorrhage [59–61]. A more recent meta-analysis in elderly AF patients which only included
324 observational studies identified similar trends with the addition that NOACs increased the risk of
325 gastro-intestinal bleeding [62]. There were no differences for this outcome in our meta-analysis, with
326 a hazard ratio of 1.04 (95% CI: 0.89-1.23) associated with NOACs compared to VKAs.

327 The fear of bleeding complications and the risk of major bleeding in patients at risk of falls or with
328 history of falls is still very low and similar for NOACs and VKAs and does not outweigh the benefits.
329 Based on these elements and the reduced risk of intracranial hemorrhage, NOACs should also be
330 considered as the first choice treatment in patients at risk of falls or with history of falls. Current
331 strategies in the AF population at risk of falls require an individualized approach that should be
332 discussed with the patients and that should consider comorbidities, costs, benefits, risks and lifestyle
333 change before anticoagulation selection to best ensure safety and compliance. These results add to
334 the existing evidence showing the improved safety and efficacy of NOACs compared to VKAs in the
335 broader AF patients and would have implications for patients, physicians and healthcare providers.

336 *Limitations and strengths*

337 This study has limitations. In particular, the included studies were not randomized; five studies were
338 retrospective, and the others were subgroup analyses of randomized clinical trials (one pre-specified

339 and four post-hoc). By design, included studies were at moderate to high risk of bias which may limit
340 the applicability of the findings as illustrated in the GRADE assessment (Table 2). To investigate the
341 effect of differential baseline prognosis between interventions, subgroup analyses were planned but
342 could not be conducted due to the limited sample size. Similarly, we could not conduct a moderator
343 analysis according to the different NOACs used but also due to the fact that some studies did not
344 specify which NOAC was considered or did not stratify the results. All studies used statistical methods
345 to adjust for differences in baseline characteristics to different extent according to the pre-specified
346 confounding domains. Nonetheless, residual confounding could not be excluded. A potential
347 limitation was confounding by indication related to previous use of VKAs. Studies considered
348 previous VKA use in their design, either by incorporating it in statistical adjustment methods (i.e.
349 multivariate cox regression, propensity score weighing) or by excluding previous VKA users. While
350 the first approach does not eliminate the possibility of residual confounding since aspects such as
351 duration of previous VKA use are not taken into account, the second may limit generalizability as
352 many NOACs users are previous VKA users [63]. There was also some variation in the definition of
353 risk of falls used by different studies. Although these definitions were similar according to our pre-
354 specified protocol, we cannot exclude the possibility that standardized population definitions would
355 have led to different results. Finally, due to the limited number of studies included, this systematic
356 literature review and meta-analysis may still be underpowered to detect small but significant
357 bleeding or thrombotic differences between VKAs and NOACs.

358 Our study has several strengths. It is the first to provide an up-to-date synthesis of the available
359 literature in a dynamically evolving field and focusing on patients at risk of falls or with history of falls
360 which have been underrepresented in the RCTs. Second, this study presents robust evidence on the
361 comparative effectiveness and safety of NOACs compared to VKAs including the use of real world
362 data which are more representative of patients being treated with anticoagulants in clinical practice.
363 Third, it uses the latest development in meta-analysis methods in the presence of dependency,
364 overcoming the limitations from the other methods suggested in Cochrane Handbook in the

365 presence of multi-arm studies [32]. These methods enable the use of all available effect sizes in the
366 analyses, so all information can be preserved and maximum statistical power is achieved [33]. Finally,
367 we used the ROBINS-I tool to evaluate the quality of the included studies, tools that enable a robust
368 assessment of the risk of different biases such as confounding or selection bias. This multilevel meta-
369 analysis highlighted the superiority of NOACs in terms of safety and efficacy compared to VKAs in AF
370 patients at risk of falls or with history of falls. Further research should be conducted to evaluate
371 which NOAC should be preferred in this patient population, using network meta-analysis methods.

372 **5. Conclusions**

373 Our systematic review and multilevel meta-analysis suggest that NOACs are reducing the risk of
374 ischemic stroke or systemic embolism (-18%) and intracranial hemorrhage (-47%) compared to VKAs
375 in patients with AF and at risk or with history of falls. There were no major differences in the risks of
376 major bleeding, ischemic stroke, hemorrhagic stroke, gastro-intestinal bleeding, cardiovascular and
377 all-cause mortality.

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380 **6. References**

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560 **7. Figures and tables**

561

562 **Fig. 1** PRISMA flow diagram of study screening and selection

563 **Fig. 2** Forest plots representing meta-analysis results comparing NOACs versus VKAs for the risk of
564 ischemic stroke and/or systemic embolism. HR: hazard ratio, CI: confidence interval, VKA: vitamin K
565 antagonist, NOAC: non-vitamin K antagonist oral anticoagulant

566 **Fig. 3** Forest plots representing meta-analysis results comparing NOACs versus VKAs for the risk of
567 intracranial hemorrhage. HR: hazard ratio, CI: confidence interval, VKA: vitamin K antagonist, NOAC:
568 non-vitamin K antagonist oral anticoagulant

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Table 1 Study characteristics

Study	Design	Population	Size	Female (%)	Age (mean or median)	HAS-BLED (mean or median)	CHA ₂ DS ₂ -VASc (mean or median)	VKA naive (%)	PAF (%)	Follow-up (mean or median - years)	Treatment group	Size	Outcomes (event rate - %/year)								
													IS and/or SE	MB	ICH	GIB	MI	IS	HS	CVM	ACM
Steffel 2016 (ENGAGE-AF-TIMI 48) [25]	Pre-specified subgroup analysis of RCT	NVAF patients at risk of falls	617	49	77	2.9	5.1	32.2	30.1	2.8	Warfarin (VKA)	307	2.85	5.55	2.08	1.46	0.88	1.78	1.01	4.99	9.98
											Edoxaban (NOAC)	310	2.81	5.43	0.33	2.91	0.89	2.17	0.38	4.18	9.27
Miao 2019 [41]	Retrospective cohort	NVAF patients at risk of falls	25,144	NR	83	NR	4	100	NR	1.4	Warfarin (VKA)	12,117	1.51	NR	0.48	NR	NR	NR	NR	NR	NR
											Rivaroxaban or Apixaban (NOAC)	13,027	1.19	NR	0.28	NR	NR	NR	NR	NR	NR
Rao 2018 (ARISTOTLE) [40]	Post-hoc subgroup analysis of RCT	NVAF patients with history of falling within 1 year	753	47	75	2.4	4.2	42.9	20.3	1.8	Warfarin (VKA)	367	1.99	5.38	1.69	NR	NR	NR	0.45	2.4	6.74
											Apixaban (NOAC)	386	1.76	4.35	0.33	NR	NR	NR	0.14	3.42	6.41
Jaspers Focks 2016 (ARISTOTLE) [42]	Post-hoc subgroup analysis of RCT	NVAF patients at risk of falls with polypharmacy (9+)	4756	46.1	71	2.3	NR	36.9	NR	1.8	Warfarin (VKA)	2380	1.79	4.21	0.97	1.08	NR	NR	NR	NR	4.85
											Apixaban (NOAC)	2376	1.35	3.55	0.28	1.23	NR	NR	NR	NR	4.55
Alexander 2019 (ARISTOTLE) [43]	Post-hoc subgroup analysis of RCT	NVAF patients at risk of falls with high multimorbidity (6+)	2222	38	74	2.4	4.9	NR	NR	1.8	Warfarin (VKA)	NR	1.80	4.88	0.84	NR	1.60	NR	0.26	NR	7.89
											Apixaban (NOAC)	NR	1.67	3.99	0.23	NR	1.14	NR	0.21	NR	6.97
Fanning 2020 [44]	Retrospective cohort	NVAF patients at risk of falls with dementia	2399	54	82	NR	NR	NR	NR	NR	Warfarin (VKA)	1386	4.82	NR	0.76	1.28	NR	2.77	NR	NR	NR
											Rivaroxaban or Apixaban (NOAC)	1013	3.9	NR	0.35	3.32	NR	2.49	NR	NR	NR
Hohmann 2019 [48]	Retrospective cohort	NVAF patients at risk of falls with frailty	36,267	61	77	NR	4.5	NR	NR	2.3	Warfarin (VKA)	NR	1.8	NR	0.94	1.71	NR	NR	NR	NR	NR
		NOAC	NR	1.78	NR	0.58	1.88	NR	NR	NR	NR	NR									
		NVAF patients at risk of falls with multimorbidity (4+)	26,410	45	77	NR	4.9	NR	NR	2.3	Warfarin (VKA)	NR	2.03	NR	1.05	2.13	NR	NR	NR	NR	NR
		NOAC	NR	2.22	NR	0.65	2.17	NR	NR	NR	NR	NR									
		NVAF patients at risk of falls with polypharmacy (7+)	33,238	51	76	NR	4.5	NR	NR	2.3	Warfarin (VKA)	NR	1.74	NR	0.94	1.78	NR	NR	NR	NR	NR
		NOAC	NR	1.74	NR	0.60	1.90	NR	NR	NR	NR	NR									
Piccini 2016 (ROCKET AF) [45]	Post-hoc subgroup analysis of RCT	NVAF patients at risk of falls with polypharmacy (10+)	1835	39	75	NR	NR	18	20	NR	Warfarin (VKA)	NR	NR	6.14	NR	NR	NR	NR	NR	NR	NR
											Rivaroxaban (NOAC)	NR	NR	6.54	NR	NR	NR	NR	NR	NR	
Lip 2021 [46]	Retrospective cohort	NVAF patients at risk of falls with frailty	NR	NR	NR	NR	NR	NR	NR	0.5-0.7	Warfarin (VKA)	34594	3.3	9.04	1.49	4.16	NR	2.35	0.80	NR	NR
											Apixaban (NOAC)	34594	2.18	6.05	0.83	2.87	NR	1.77	0.34	NR	NR
											Warfarin (VKA)	9263	3.06	8.89	1.42	4.29	NR	2.04	0.81	NR	NR
											Dagibatan (NOAC)	9263	2.6	7.07	0.65	3.99	NR	2.2	0.28	NR	NR
											Warfarin (VKA)	39898	3.13	8.88	1.45	4.22	NR	2.22	0.78	NR	NR
											Rivaroxaban (NOAC)	39898	2.5	10.24	1.03	5.63	NR	1.85	0.53	NR	NR
Martinez 2018 [47]	Retrospective cohort	NVAF patients at risk of falls with frailty	19077	NR	NR	NR	NR	NR	NR	0.9-1.8	Warfarin (VKA)	1392	2.15	4.41	0.37	3.09	NR	2.0	0.15	NR	NR
											Apixaban (NOAC)	1392	1.68	3.11	0.35	2.33	NR	1.4	0.28	NR	NR
											Warfarin (VKA)	1350	2.2	4.44	0.59	3.31	NR	1.93	0.32	NR	NR
											Dagibatan (NOAC)	1350	2.06	3.82	0.10	3.10	NR	1.73	0.10	NR	NR
											Warfarin (VKA)	2635	2.61	4.01	0.60	2.70	NR	2.18	0.36	NR	NR
											Rivaroxaban (NOAC)	2635	1.78	4.13	0.29	3.41	NR	1.51	0.26	NR	NR

572 RCT: randomized clinical trial, NVAf: non-valvular atrial fibrillation, NR: not reported, VKA: vitamin K antagonist, NOAC: non-vitamin K antagonist oral anticoagulant, QD:
573 once a day, BID: twice a day, PAF: paroxysmal atrial fibrillation, IS: ischemic stroke, SE: systemic embolism, MB; major bleeding, GIB: gastro-intestinal bleeding, ICH:
574 intracranial hemorrhage, MI: myocardial infarction, HS: hemorrhagic stroke, CVM: cardiovascular mortality, ACM: all-cause mortality

575 Table 2 Summary of findings according to GRADE

NOACs compared to VKAs for the management of AF patient at risk (or with history) of falls

Patient or population: AF patient at risk (or with history) of falls

Comparison: VKAs

Intervention: NOACs

Outcomes	Anticipated absolute effects* (95% CI)		Relative effect (95% CI)	No of participants (studies)	Certainty of the evidence (GRADE)
	Risk with VKAs	Risk with NOACs			
Ischemic stroke and/or systemic embolism follow-up: range 0.5 years to 2.8 years	Moderate 24 per 1 000		HR 0.82 (0.69 to 0.98)	Range from 617 to 79,796 (15 effect sizes, 7 non-randomized studies)	⊕○○○ Very low ^{a,b}
Intracranial hemorrhage follow-up: range 0.5 years to 2.8 years	Moderate 10 per 1 000		HR 0.53 (0.40 to 0.71)	Range from 617 to 79,796 (15 effect sizes, 7 non-randomized studies)	⊕○○○ Very low ^{a,c}
Major bleeding follow-up: range 0.5 years to 2.8 years	Moderate 60 per 1 000		HR 0.88 (0.74 to 1.04)	Range from 617 to 79,796 (11 effect sizes, 5 non-randomized studies)	⊕○○○ Very low ^a
Ischemic stroke follow-up: range 0.5 years to 2.8 years	Moderate 22 per 1 000		HR 0.87 (0.60 to 1.28)	Range from 617 to 79,796 (8 effect sizes, 4 non-randomized studies)	⊕○○○ Very low ^{b,d,f}
Hemorrhagic stroke follow-up: range 0.5 years to 2.8 years	Moderate 6 per 1 000		HR 0.51 (0.24 to 1.10)	Range from 617 to 79,796 (9 effect sizes, 4 non-randomized studies)	⊕○○○ Very low ^{a,d}
Gastro-intestinal bleeding follow-up: range 0.5 years to 2.8 years	Moderate 26 per 1 000		HR 1.04 (0.89 to 1.23)	Range from 617 to 79,796 (12 effect sizes, 6 non-randomized studies)	⊕○○○ Very low ^{a,b,d}
Myocardial infarction follow-up: range 1.8 years to 2.8 years	Moderate 9 per 1 000		HR 0.76 (0.47 to 1.24)	Range from 617 to 2222 (2 non-randomized studies)	⊕⊕○○ Low ^{d,e}
Cardiovascular mortality follow-up: range 0.5 years to 2.8 years	Moderate 37 per 1 000		HR 1.04 (0.61 to 1.75)	1370 (2 non-randomized studies)	⊕⊕○○ Low ^{d,e}
All-cause mortality follow-up: range 1.4 years to 2.8 years	Moderate 74 per 1 000		HR 1.23 (0.35 to 4.29)	Range from 617 to 4756 (5 effect sizes, 3 non-randomized studies)	⊕○○○ Very low ^{b,d,e}

*The risk in the intervention group (and its 95% confidence interval) is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% CI).

GRADE Working Group grades of evidence

High certainty: we are very confident that the true effect lies close to that of the estimate of the effect.

Moderate certainty: we are moderately confident in the effect estimate: the true effect is likely to be close to the estimate of the effect, but there is a possibility that it is substantially different.

Low certainty: our confidence in the effect estimate is limited: the true effect may be substantially different from the estimate of the effect.

Very low certainty: we have very little confidence in the effect estimate: the true effect is likely to be substantially different from the estimate of effect.

576 ^aThere was serious bias due to confounding and deviation from the intended intervention. ^bThe
577 heterogeneity was substantial. ^cAsymmetrical Funnel plot and significant Egger's regression test.
578 ^dBoundaries of the CI cross the clinical decision threshold. ^eThere was serious bias due to confounding.
579 ^fThere was serious bias due to deviation from the intended intervention. HR: hazard ratio, CI: confidence
580 interval VKA: vitamin K antagonist, NOAC: non-vitamin K antagonist oral anticoagulant, AF: atrial
581 fibrillation