



## LJMU Research Online

**Lolli, L, Bahr, R, Weston, M, Whiteley, R, Tabben, M, Bonanno, D, Gregson, W, Chamari, K, Di Salvo, V and van Dyk, N**

**No association between perceived exertion and session duration with hamstring injury occurrence in professional football**

<http://researchonline.ljmu.ac.uk/id/eprint/12883/>

### Article

**Citation** (please note it is advisable to refer to the publisher's version if you intend to cite from this work)

**Lolli, L, Bahr, R, Weston, M, Whiteley, R, Tabben, M, Bonanno, D, Gregson, W, Chamari, K, Di Salvo, V and van Dyk, N (2019) No association between perceived exertion and session duration with hamstring injury occurrence in professional football. Scandinavian Journal of Medicine & Science in**

LJMU has developed [LJMU Research Online](#) for users to access the research output of the University more effectively. Copyright © and Moral Rights for the papers on this site are retained by the individual authors and/or other copyright owners. Users may download and/or print one copy of any article(s) in LJMU Research Online to facilitate their private study or for non-commercial research. You may not engage in further distribution of the material or use it for any profit-making activities or any commercial gain.

The version presented here may differ from the published version or from the version of the record. Please see the repository URL above for details on accessing the published version and note that access may require a subscription.

For more information please contact [researchonline@ljmu.ac.uk](mailto:researchonline@ljmu.ac.uk)

<http://researchonline.ljmu.ac.uk/>

1 **No association between measures of perceived exertion and session duration with**  
2 **hamstring injury occurrence in professional football**

3  
4  
5 **Authors:** Lorenzo Lolli<sup>1,6</sup>, Roald Bahr<sup>2,3</sup>, Matthew Weston<sup>1,4</sup>, Rodney Whiteley<sup>2</sup>  
6 Montassar Tabben<sup>5</sup>, Daniele Bonanno<sup>1</sup>, Warren Gregson<sup>1,6</sup>, Karim Chamari<sup>5</sup>, Valter Di  
7 Salvo<sup>1,7</sup>, Nicol van Dyk<sup>2</sup>

8  
9  
10  
11 <sup>1</sup> *Aspire Academy, Football Performance & Science Department, Doha, Qatar*

12 <sup>2</sup> *Sport Medicine Department, Aspetar, Qatar Orthopaedic and Sports Medicine Hospital,*  
13 *Doha, Qatar.*

14 <sup>3</sup> *Oslo Sports Trauma Research Center, Norwegian School of Sport Sciences, Oslo, Norway*

15 <sup>4</sup> *School of Health and Social Care, Teesside University, Middlesbrough, UK*

16 <sup>5</sup> *Athlete Health and Performance Research Centre, Aspetar, Qatar Orthopaedic and Sports*  
17 *Medicine Hospital, Doha, Qatar.*

18 <sup>6</sup> *Football Exchange, Research Institute of Sport Sciences, Liverpool John Moores*  
19 *University, Liverpool, UK*

20 <sup>7</sup> *Department of Movement, Human and Health Sciences, University of Rome “Foro*  
21 *Italico”, Rome, Italy*

22  
23 **Correspondence:**

24 Lorenzo Lolli, Aspire Academy & Qatar FA, PO Box 22287, Doha, Qatar.

25 e-mail: [Lorenzo.Lolli@aspire.qa](mailto:Lorenzo.Lolli@aspire.qa)

26  
27  
28  
29  
30  
31  
32  
33  
34  
35  
36  
37  
38  
39  
40  
41  
42

43 **Abstract**

44 Training and competition loads have emerged as modifiable composite risk factors of non-  
45 contact injury. Hamstring strains are the most common injuries in football with substantial  
46 burden on the individual player and club. Nevertheless, robust evidence of a consistent load-  
47 hamstring injury relationship in professional football is lacking. Using available data from the  
48 Qatar Stars League over three competitive seasons, this study investigated the separate and  
49 combined effects of perceived exertion and session duration on hamstring injury occurrence in  
50 a sample of 30 outfield football players. Load variables were calculated into 7-day, 14-day, 21-  
51 day, 28-day periods of data, and week-to-week changes for average ratings of perceived  
52 exertion (RPE; au) score and session-RPE (s-RPE; session-duration  $\times$  score), plus the  
53 cumulative training and match minutes and s-RPE, respectively. Conditional logistic  
54 regression models estimated load-injury relationships per 2-within-subject standard deviation  
55 increments in each candidate variable. Associations were declared practically important based  
56 on the location of the confidence interval in relation to thresholds of 0.90 and 1.11 defining  
57 small beneficial and harmful effects, respectively. The uncertainty for the corrected odds ratios  
58 show that typically high within-subject increments in each candidate variable were not  
59 practically important for training- and match-related hamstring injury (95% confidence  
60 intervals range: 0.85 to 1.16). We found limited exploratory evidence regarding the value of  
61 measures of perceived exertion and session duration as aetiological factors of hamstring injury  
62 in Middle-East professional football. Monitoring remains valuable to inform player load  
63 management strategies, but our exploratory findings suggest its role for type-specific injury  
64 risk determination appears empirically unsupported.

65 **Keywords:** hamstrings, load, perceived exertion, RPE, muscle injury, risk factors

66  
67  
68  
69  
70  
71  
72  
73  
74  
75  
76  
77  
78  
79  
80  
81  
82  
83  
84  
85  
86  
87  
88  
89  
90  
91

## 92 Introduction

93 Hamstring injury is the most common type of non-contact muscle injury in elite football, with  
94 one injury every 1000 h of play leading to 19 days lost from training and match-play.<sup>1,2</sup> Until  
95 2015, hamstring injury incidence increased annually by 2.3%, with an economic burden of  
96 £74.4 million in elite European football.<sup>3-5</sup> Also, the risk of re-injuries is substantial and non-  
97 contact injuries can impact team performance negatively.<sup>6</sup>

98 Although many risk factors for hamstring injury have been investigated [i.e., strength,  
99 flexibility, and previous injury],<sup>7,8</sup> no work has evaluated the contribution of training and  
100 competition loads on hamstring injury risk. This is somewhat surprising given the increasing  
101 load demands<sup>9</sup> and congested fixtures<sup>10</sup> in elite football and a primary purpose of monitoring  
102 training loads in elite football is injury reduction.<sup>10</sup> From an applied standpoint, a clear  
103 understanding of the association between load and non-contact hamstring injury is an  
104 important, yet preliminary, step in the process for developing interventions to optimise  
105 performance and maximise player availability.

106 Previous examinations of the load-injury relationship in elite football players have a number  
107 of limitations, including the injury groups used as outcome measures, the load metrics used as  
108 exposure measures and the study designs. First, studies have combined a range of different  
109 injury types as outcome measure and it is unlikely that the load-injury relationship is the same  
110 for different acute injury types (e.g., hamstring strains and ankle sprains) or overuse injuries  
111 (e.g., metatarsal stress fractures and patellar tendinopathy). No study has yet examined the  
112 relationship between a single injury type and load. Second, studies have calculated acute and  
113 chronic external and internal loads represented by prior 7-, 14-, 21-, and 28-day loads, week-  
114 to-week changes, and the acute:chronic workload ratio (ACWR), with inconsistent findings.<sup>11-</sup>  
115 <sup>16</sup> Despite inherent limitations of this ratio for applied and medical purposes,<sup>17,18</sup> recent studies  
116 in football have examined associations between typically high ACWR values and increased  
117 non-contact injury risk.<sup>12,13,16</sup> Furthermore, transforming continuous measures of load into  
118 categorical variables (e.g., high, moderate, low) involves a loss of statistical power, increased  
119 Type I error rates, and an underestimation of the variation in the outcome of interest.<sup>19</sup> Third,  
120 previous research has compared the load pattern of injured players to that of their uninjured  
121 teammates.<sup>12-16,20</sup> It seems more appropriate to compare injured players to themselves, i.e.,  
122 whether the load pattern preceding injury differs from their usual load. Finally, previous  
123 investigations used a composite measure of internal load that combines training and  
124 competition duration with perceived exertion (session-RPE, s-RPE).<sup>12,13,15,16</sup> While this  
125 approach is useful for quantifying weekly and training phase load, a specific breakdown is  
126 unclear as the score neglects quantification of intensity and duration in isolation, both of which  
127 are important for effective training planning.<sup>21</sup>

128 We therefore designed the present study to examine the effect of load on acute hamstring injury  
129 occurrence, the most important type of injury in professional football, using continuous  
130 measures of perceived intensity and session duration and adopting the normal load pattern of  
131 injured players as our control comparison.

## 132 Methods

### 133 Participants

134 Study participants included outfield professional football players competing in the Qatar Stars  
135 League (QSL) over three seasons (May 2015 to February 2018). A complete overview of the

136 injury surveillance database assessment process and the final number of observations included  
137 in the study is illustrated in Figure 1. The Anti-Doping Laboratory Institutional Review Board,  
138 Qatar (protocol number: E2017000252) granted ethics approval.

### 139 **Aspetar Injury and Illness Surveillance Programme**

140 Injury information was retrieved as part of the medical services provided to all participating  
141 QSL teams by the National Sports Medicine Programme within the Aspetar Orthopaedic and  
142 Sports Medicine Hospital. This centralized system with a focal point for the medical care of  
143 each club competing in the QSL allowed for standardization of the Aspetar Injury and Illness  
144 Surveillance Programme.<sup>22</sup> This programme includes prospective injury registration from all  
145 QSL teams. Injury data were collected prospectively, with monthly reporting and regular  
146 communication with the responsible team physician/physiotherapist to encourage timely and  
147 accurate reporting. As detailed previously,<sup>7,8</sup> a traumatic hamstring injury (i.e., sudden onset  
148 injury) was defined as acute pain in the posterior thigh that occurred during training or match  
149 play and resulted in immediate termination of all activity and a subsequent inability to  
150 participate in the next training session or match. These injuries were confirmed through a  
151 clinical examination (identifying pain on palpation, pain with isometric contraction, and pain  
152 with muscle lengthening) by the team physician. If indicated, the clinical diagnosis was  
153 supported by ultrasonography and magnetic resonance imaging at the study centre. Figure 1  
154 depicts the inclusion methodology during the three study seasons. Only injuries that resulted  
155 in more than three days of absence were included in this study, calculated from the date of  
156 injury to the date of the player's return to full unrestricted participation in team training and  
157 availability for match selection. Recurrent hamstring injuries were excluded from the primary  
158 analysis.

### 159 **Load monitoring**

160 Training and match loads were quantified as session duration (minutes) and RPE. Players rated  
161 the global intensity of all sessions and matches using level-anchored semi-ratio CR-10 Borg  
162 scale (Borg CR10<sup>®</sup>).<sup>23</sup> Science and/or medicine staff collected RPE ~30 min after completion  
163 of the session/match.

### 164 **Calculation of load variables**

165 The study sample included only players with a minimum of two-months of complete  
166 measurements after the first official match of the season, and players with insufficient in-season  
167 data precluding the calculation of the predefined time periods free from the influence of the  
168 pre-season data were excluded from the analyses (Figure 1). Where available, given the  
169 retrospective nature of the present study, the injury load day value was included in the  
170 calculation. If not recorded, the load calculation considered the observation of the day prior to  
171 hamstring injury occurrence. In the case of missing values for the load variable with complete  
172 outcome data information, the sample-based session-specific median value for either training  
173 or match-play was assigned for missing load observations in the available data set (9.6%).  
174 Table 1 provides a detailed illustration of an example dataset of one player showing the data  
175 structure for performance and injury data required for this study. We calculated the following  
176 exposure variables: i) average RPE score, ii) average s-RPE (session duration × score), iii)  
177 cumulative exposure in minutes, and iv) cumulative s-RPE calculated over 7-day, 14-day, 21-  
178 day, and 28-day periods. In addition to this, week-to-changes for cumulative duration in  
179 minutes and s-RPE were derived.<sup>16</sup> These data were, therefore, calculated into the predefined  
180 load periods in which the injury (i) occurred and (ii) did not occur (Table 1). As an example,

181 for illustrating how each variable was calculated, Figure 2 shows data for a player's 7-day  
182 average s-RPE leading into an injury. Data for each variable were considered only for the  
183 season in which an injury occurred.

## 184 **Statistical analysis**

185 The number of time-loss days for hamstring injury are summarised as median and interquartile  
186 range (IQR). Conditional fixed-effects logistic regression analyses estimated the odds of  
187 experiencing a hamstring injury based on the comparison of players' injury load data versus  
188 control data in which an injury did not occur using the *survival* package. This procedure is  
189 different from the conventional logistic regression modelling, whereby the calculation of the  
190 conditional likelihood involved the analysis of load data with player identity as a cluster factor  
191 in the model to account for the within-subject association between the examined  
192 observations.<sup>24</sup> The relationship between each variable with hamstring injury was examined  
193 for the first event only. To examine the association between training load and hamstring injury  
194 occurrence, odds ratios (OR) were derived for a 2-within-player SD increment in each  
195 variable,<sup>25</sup> representing the effect of a typically high versus a typically low value.<sup>26</sup> A within-  
196 player SD of the variables was calculated as the square root of the residual mean square.<sup>27</sup>  
197 Thresholds of 0.9, 0.7, 0.5, 0.3 and 0.1 and their reciprocals 1.11, 1.43, 2.0, 3.3 and 10 defined  
198 small, moderate, large, very large and extremely large beneficial and harmful effects,  
199 respectively.<sup>26</sup> Retrospective design analyses assessed Type M error rates for the point  
200 estimates and sampling uncertainty of the observed effects.<sup>28</sup> This approach provides an  
201 objective quantification of the degree of overestimation of an observed effect estimate relative  
202 to the magnitude of the true underlying population effect given the data.<sup>28</sup> Corrected ORs were  
203 obtained by dividing the natural logarithm of the estimated OR by the respective magnitude of  
204 exaggeration or Type M error relative to a targeted small increase or reduction in the odds of  
205 injury of  $\ln\text{OR} = \pm |0.105360515657826|$ . In the absence of an established anchor defining a  
206 practically important increase or reduction in the odds of sustaining a hamstring injury, we  
207 considered a 10% lower (OR = 0.90) or a 11% higher (OR = 1.11) odds of clinical event as  
208 substantially beneficial and substantially harmful effects, respectively.<sup>26</sup> Associations were  
209 therefore declared practically important based on the location of the confidence interval for the  
210 estimated true ORs to these thresholds.

211 Since this is the first study to examine the relationship between load and hamstring injury in  
212 football, a formal a priori sample size estimation was not possible using existing studies as per  
213 the TRIPOD (Transparent Reporting of a multivariable prediction model for Individual  
214 Prognosis Or Diagnosis) statement 22-item checklist.<sup>29</sup> Accordingly, to inform the design of  
215 future studies,<sup>30</sup> Cox-Snell pseudo- $R^2$  ( $R^2_{CS}$ ) statistics were reported as measures of model  
216 overall performance.<sup>31</sup> Outcome statistics are reported as point estimates and 95% confidence  
217 intervals (CI). Statistical analyses were performed using R (version 3.5.1, R Foundation for  
218 Statistical Computing, Vienna, Austria).

## 219 **Results**

220 Overall, 30 outfield football players with valid physical load and hamstring injury data were  
221 eligible for this study (Figure 1). A total of 145 injuries were excluded from the analysis; 3  
222 were recurrent injuries, 18 due to reporting error and 124 due to insufficient exposure data. The  
223 median time-loss days for hamstring injury was 18 (IQR, 13 to 25). Irrespective of different  
224 approaches for the calculation of load data over predefined time periods, the corrected odds of  
225 hamstring injury in the average RPE score, average s-RPE, cumulative duration in minutes,

226 and cumulative s-RPE for all the physical load periods were not practically important (Table  
227 2).

## 228 **Discussion**

229 This is the first study examining the relationship of match and training load with acute  
230 hamstring injuries in professional football. Using a research design and methodological  
231 framework addressing common shortcomings in the current literature, we did not find any  
232 practically relevant association between measures of perceived exertion and session duration  
233 with hamstring injury occurrence in professional football players.

234 Load monitoring is critical to inform medical and performance staff strategies.<sup>32</sup> Previous  
235 investigations into associations of load with non-contact injury occurrence in football  
236 examined the prognostic value of composite measures of external and internal load as potential  
237 risk factors yielding unclear and inconsistent findings.<sup>11-16</sup> However, these studies were not  
238 without methodological shortcomings, most notably the use of ratio indices, multiple load time  
239 bins analysed as categorical variables, and a composite score.<sup>18,19,21</sup> Additionally, the failure of  
240 researchers to distinguish the specific nature of an event within the spectrum of acute or overuse  
241 injuries represents an additional limitation substantiating the limited practical utility of load-  
242 injury studies in the available literature.<sup>11-16</sup> The lack of a clear differentiation between injury  
243 types as outcome measures implies that the load-injury relationship is assumed to be same  
244 within the spectrum of acute or overuse injuries, which appears implausible on clinical  
245 grounds. Therefore, also depending on which external or internal load measure is selected as  
246 exposure variable, we maintain that a precise definition of the injury type is fundamental to  
247 provide information about the odds or risk of type-specific injury to inform medical and  
248 performance staff meaningfully.

249 From applied and clinical perspectives, the present study advances our understanding of the  
250 load-hamstring injury relationship in professional football. The notion of physical load  
251 involves an understanding of the interplay between intensity, volume, and frequency to  
252 determine training outcome,<sup>23</sup> yet this is underappreciated in the load-injury literature. While  
253 technological advances now permit a detailed measurement of player external load,<sup>33</sup> when  
254 compared with s-RPE measures, quantification of external load via global positioning system  
255 (GPS) fails to represent the actual physiological stress imposed upon players.<sup>33</sup> Despite being  
256 widely adopted in this context, s-RPE is not without limitation as a global measure of effort  
257 perception. It might underrepresent the stochastic demands of football<sup>23</sup> and obfuscate the  
258 separate effects and contribution of intensity and duration on the training process.<sup>21</sup>

259 Previous examinations of the load-injury relationship in elite football players have reported  
260 inconsistent findings regarding the association with loading derived from various time  
261 windows.<sup>13,15,16</sup> Irrespective of the use of different time windows and alternative approaches  
262 for the calculation of training and competition loads in the present study, we did not find any  
263 effect of separate and combined measures of intensity and duration on hamstring injury  
264 occurrence were not practically important (Table 2). From a real-world perspective, current  
265 match schedule informs the training plan and weekly schedules (i.e., 7-day) are designed to  
266 ensure players are match ready.<sup>10,34</sup> In this context, 7-day and 28-day periods would represent  
267 logical and practical units to define short- and long-term physical loads.<sup>10</sup> The use of multiple  
268 time periods to determine physical loads likely adds a further layer of unnecessary complexity,  
269 and it might have contributed to the inconsistency of studies in football.

270 The methodological flaws in the current field of research<sup>11-16,32</sup> should be considered when  
271 interpreting the available data. In particular, the conceptual and statistical flaws of  
272 indiscriminate categorisation of continuous variables for prognostic model development are  
273 well-established.<sup>19</sup> Recently, the pitfalls of indiscriminate discretization were illustrated in the  
274 case of regression modelling strategies involving measures of physical load entered as  
275 categorical variables.<sup>19</sup> With this in mind, using more appropriate conditional modelling  
276 strategies<sup>35</sup> given the present study design, we estimated the effects per 2-within-player SD  
277 increment in the exposure<sup>25,36</sup> and therefore avoided inappropriate discrete approaches as  
278 illustrated in a previous study.<sup>20</sup> Despite the available approaches for modelling training and  
279 competition loads,<sup>19,20</sup> estimation of the within-player variance may be a simpler and valid  
280 approach to determine reference ranges for player load monitoring and guide  
281 interpretations.<sup>27,36</sup> Although variance is generally used to describe measurement error,  
282 estimation of the within-player variability might represent a valuable alternative to facilitate  
283 the longitudinal tracking of training and competition loads over time both for research and  
284 applied purposes. The present study is the first to investigate the load-injury relationship in  
285 football using a within-subject analysis. As illustrated in Figure 1, we lost over 80% of the  
286 players eligible for this study to follow up and this was due to a lack of accurate data collection,  
287 or insufficient data to perform the appropriate analysis. From applied and clinical perspectives,  
288 this highlights the challenges in this type of data collection.

## 289 **Limitations**

290 Given the novelty of our study, a formal a priori sample size estimation informed by the  
291 precision of coefficient estimates<sup>37</sup> or relevant model statistics<sup>31</sup> from any existing study could  
292 not be performed. Nevertheless, recent advances in the procedures for determining minimum  
293 sample size now permit a robust appraisal of the sample size requirements based on pseudo-R<sup>2</sup>  
294 statistics.<sup>30</sup> Therefore, we reported the recommended statistics<sup>30</sup> which can be used by  
295 researchers and clinicians to inform sample size estimation for future investigations in this field  
296 (Table 2). For example, in the case of the model with the 28-day cumulative session duration,  
297 assuming a population outcome prevalence of 0.309<sup>7</sup> and using the R<sup>2</sup><sub>CS</sub> value of 0.074 in the  
298 equation indicate a minimum sample size requirement of 329, 583, 1166 players for the  
299 development of new models with one, five, and ten load-related candidate predictor  
300 parameters, respectively.

301 In the present study, internal load was quantified using RPE, which represents a global measure  
302 of session intensity. While this measure is practical, it fails to capture the whole range of  
303 football-related perceptual sensations.<sup>38</sup> Similar to the quantification of the physical  
304 performance demands based on relevant measures of external load,<sup>39</sup> the use of differential  
305 RPE would represent a valuable alternative here as it provides greater precision in scaling  
306 psychophysiological signals during training and match-play and therefore enhances  
307 understanding of how different dimensions of exertion contribute to overall physical exertion.<sup>38</sup>  
308 From a medical perspective, differential RPE may also be of particular relevance for the study  
309 of type-specific soft-tissue injuries aetiology (e.g., peripherally dominated ratings on the Borg  
310 scale).<sup>38</sup>

311 A clear distinction between match and training loads might also be necessary. For example, in-  
312 season loads are substantially lower in training than during official match-play<sup>40</sup> and the  
313 occurrence of hamstring injuries is higher during match-play than training.<sup>1</sup> Therefore,  
314 competition load could determine higher risk for non-contact injuries, so investigating how  
315 different physical efforts undertaken during match-play contribute to hamstring strains appears  
316 warranted. Finally, the potential homogeneity of the present study cohort, representative of



317 mainly Middle East professional football players, training culture, and specific regional  
318 climatic conditions are all factors limiting the generalisability of our study findings to other  
319 contexts.

## 320 **Perspective**

321 We found no preliminary evidence of associations between hamstring injuries and measures of  
322 perceived exertion intensity or session duration that may suggest a role in the aetiology of this  
323 type of injury. While longitudinal tracking of changes in training and competition loads  
324 remains important for informing the player management process, our exploratory study  
325 suggests that the use of separate or combined measures of perceived exertion and session  
326 duration in examining the load-hamstring injury relationship is not empirically supported. For  
327 the first time, given the novelty of our investigation, we also provide distinct  $R^2_{CS}$  estimates  
328 which are anticipated to serve as a guide to inform sample size calculations in future studies  
329 on load and hamstring injury occurrence in professional football.

## 330 **References**

- 331 1. Ekstrand J, Walden M, Hagglund M. Hamstring injuries have increased by 4%  
332 annually in men's professional football, since 2001: a 13-year longitudinal analysis of  
333 the UEFA Elite Club injury study. *Br J Sports Med.* 2016;50(12):731-737.
- 334 2. Bahr R, Clarsen B, Ekstrand J. Why we should focus on the burden of injuries and  
335 illnesses, not just their incidence. *Br J Sports Med.* 2018;52(16):1018-1021.
- 336 3. Ekstrand J. Keeping your top players on the pitch: the key to football medicine at a  
337 professional level. *Br J Sports Med.* 2013(47):723-724.
- 338 4. de Visser HM, Reijman M, Heijboer MP, Bos PK. Risk factors of recurrent hamstring  
339 injuries: a systematic review. *Br J Sports Med.* 2012;46(2):124-130.
- 340 5. Wangensteen A, Tol JL, Witvrouw E, et al. Hamstring reinjuries occur at the same  
341 location and early after return to sport: a descriptive study of MRI-confirmed  
342 reinjuries. *Am J Sports Med.* 2016;44(8):2112-2121.
- 343 6. Hagglund M, Walden M, Magnusson H, Kristenson K, Bengtsson H, Ekstrand J.  
344 Injuries affect team performance negatively in professional football: an 11-year  
345 follow-up of the UEFA Champions League injury study. *Br J Sports Med.*  
346 2013;47(12):738-742.
- 347 7. van Dyk N, Bahr R, Whiteley R, et al. Hamstring and quadriceps isokinetic strength  
348 deficits are weak risk factors for hamstring strain injuries: a 4-year cohort study. *Am J*  
349 *Sports Med.* 2016;44(7):1789-1795.
- 350 8. van Dyk N, Farooq A, Bahr R, Witvrouw E. Hamstring and ankle flexibility deficits  
351 are weak risk factors for hamstring Injury in professional soccer players: a  
352 prospective cohort study of 438 players including 78 injuries. *Am J Sports Med.*  
353 2018;46(9):2203-2210.
- 354 9. Barnes C, Archer DT, Hogg B, Bush M, Bradley PS. The evolution of physical and  
355 technical performance parameters in the English Premier League. *Int J Sports Med.*  
356 2014;35(13):1095-1100.
- 357 10. Weston M. Training load monitoring in elite English soccer: a comparison of  
358 practices and perceptions between coaches and practitioners. *Sci Med Footb.* 2018.
- 359 11. Bowen L, Gross AS, Gimpel M, Bruce-Low S, Li FX. Spikes in acute:chronic  
360 workload ratio (ACWR) associated with a 5-7 times greater injury rate in English  
361 Premier League football players: a comprehensive 3-year study. *Br J Sports Med.*  
362 2019.

- 363 12. Delecroix B, McCall A, Dawson B, Berthoin S, Dupont G. Workload and non-contact  
364 injury incidence in elite football players competing in European leagues. *Eur J Sport*  
365 *Sci.* 2018;18(9):1280-1287.
- 366 13. Fanchini M, Rampinini E, Riggio M, Coutts AJ, Pecci C, McCall A. Despite  
367 association, the acute:chronic work load ratio does not predict non-contact injury in  
368 elite footballers. *Sci Med Footb.* 2018:108-114.
- 369 14. Jaspers A, Kuyvenhoven JP, Staes F, Frencken WGP, Helsen WF, Brink MS.  
370 Examination of the external and internal load indicators' association with overuse  
371 injuries in professional soccer players. *J Sci Med Sport.* 2018;21(6):579-585.
- 372 15. Malone S, Owen A, Newton M, Mendes B, Collins KD, Gabbett TJ. The acute:chronic  
373 workload ratio in relation to injury risk in professional soccer. *J Sci Med Sport.* 2016.
- 374 16. McCall A, Dupont G, Ekstrand J. Internal workload and non-contact injury: a one-  
375 season study of five teams from the UEFA Elite Club Injury Study. *Br J Sports Med.*  
376 2018.
- 377 17. Lolli L, Batterham AM, Hawkins R, et al. Mathematical coupling causes spurious  
378 correlation within the conventional acute-to-chronic workload ratio calculations. *Br J*  
379 *Sports Med.* 2017.
- 380 18. Lolli L, Batterham AM, Hawkins R, et al. The acute-to-chronic workload ratio: an  
381 inaccurate scaling index for an unnecessary normalisation process? *Br J Sports Med.*  
382 2018.
- 383 19. Carey DL, Crossley KM, Whiteley R, et al. Modelling training loads and injuries: the  
384 dangers of discretization. *Med Sci Sports Exerc.* 2018.
- 385 20. Windt J, Arden CL, Gabbett TJ, et al. Getting the most out of intensive longitudinal  
386 data: a methodological review of workload-injury studies. *BMJ Open.*  
387 2018;8(10):e022626.
- 388 21. Juhari F, Ritchie D, O'Connor F, et al. The quantification of within-week session  
389 intensity, duration, and intensity distribution across a season in Australian football  
390 using the session rating of perceived exertion method. *Int J Sports Physiol Perform.*  
391 2018;13(7):940-946.
- 392 22. Bakken A, Targett S, Bere T, et al. Health conditions detected in a comprehensive  
393 periodic health evaluation of 558 professional football players. *Br J Sports Med.*  
394 2016;50(18):1142-1150.
- 395 23. McLaren SJ, Macpherson TW, Coutts AJ, Hurst C, Spears IR, Weston M. The  
396 relationships between internal and external measures of training load and intensity in  
397 team sports: a meta-analysis. *Sports Med.* 2018;48(3):641-658.
- 398 24. Connolly MA, Liang KY. Conditional logistic regression models for correlated binary  
399 data. *Biometrika.* 1988;75(3):501-506.
- 400 25. Gelman A. Scaling regression inputs by dividing by two standard deviations. *Stat*  
401 *Med.* 2008;27(15):2865-2873.
- 402 26. Sharma J, Weston M, Batterham AM, Spears IR. Gait retraining and incidence of  
403 medial tibial stress syndrome in army recruits. *Med Sci Sports Exerc.*  
404 2014;46(9):1684-1692.
- 405 27. Bland JM, Altman DG. Measurement error. *BMJ.* 1996;313(7059):744.
- 406 28. Gelman A, Carlin J. Beyond power calculations: assessing type S (sign) and type M  
407 (magnitude) errors. *Perspect Psychol Sci.* 2014;9(6):641-651.
- 408 29. Moons KG, Altman DG, Reitsma JB, et al. Transparent Reporting of a multivariable  
409 prediction model for Individual Prognosis or Diagnosis (TRIPOD): explanation and  
410 elaboration. *Ann Intern Med.* 2015;162(1):W1-73.

- 411 30. Riley RD, Snell KI, Ensor J, et al. Minimum sample size for developing a  
412 multivariable prediction model: PART II - binary and time-to-event outcomes. *Stat*  
413 *Med.* 2018.
- 414 31. Steyerberg EW, Vickers AJ, Cook NR, et al. Assessing the performance of prediction  
415 models: a framework for traditional and novel measures. *Epidemiology.*  
416 2010;21(1):128-138.
- 417 32. Bourdon PC, Cardinale M, Murray A, et al. Monitoring athlete training loads:  
418 consensus statement. *Int J Sports Physiol Perform.* 2017;12(Suppl 2):S2161-S2170.
- 419 33. Weston M. Difficulties in determining the dose-response nature of competitive soccer  
420 matches. *J Athl Enhancement.* 2013;2(1).
- 421 34. Thorpe RT, Atkinson G, Drust B, Gregson W. Monitoring fatigue status in elite team-  
422 sport athletes: implications for practice. *Int J Sports Physiol Perform.* 2017;12(Suppl  
423 2):S227-s234.
- 424 35. Hu FB, Goldberg J, Hedeker D, Flay BR, Pentz MA. Comparison of population-  
425 averaged and subject-specific approaches for analyzing repeated binary outcomes. *Am*  
426 *J Epidemiol.* 1998;147(7):694-703.
- 427 36. Carroll RJ. Variances are not always nuisance parameters. *Biometrics.*  
428 2003;59(2):211-220.
- 429 37. Borenstein M. Planning for precision in survival studies. *J Clin Epidemiol.*  
430 1994;47(11):1277-1285.
- 431 38. Weston M, Siegler J, Bahnert A, McBrien J, Lovell R. The application of differential  
432 ratings of perceived exertion to Australian Football League matches. *J Sci Med Sport.*  
433 2015;18(6):704-708.
- 434 39. Gregson W, Di Salvo V, Varley M, et al. Harmful association of sprinting with  
435 muscle injury occurrence in professional soccer match-play: a two-season, league  
436 wide exploratory investigation from the Qatar Stars League. *J Sci Med Sport.* 2019.
- 437 40. Stevens TGA, de Ruiter CJ, Twisk JWR, Savelsbergh GJP, Beek PJ. Quantification of  
438 in-season training load relative to match load in professional Dutch Eredivisie football  
439 players. *Sci Med Footb.* 2017;1(2):117-125.

440  
441  
442  
443  
444  
445  
446  
447  
448  
449  
450  
451  
452  
453  
454  
455  
456  
457  
458  
459  
460

461 **Figure legends**

462

463 Figure 1. Flow diagram of the hamstring injury eligibility assessment process.

464

465 Figure 2. Descriptive characteristics a player's 7-day average s-RPE leading into an injury as  
466 an illustrative example of variable calculation. Black dots identify the observed values and the  
467 grey-shaded area defines the 95% confidence interval for the conditional-smoothed mean over  
468 the player's observational period.

469

470

471 **Table legends**

472

473 Table 1. Structure of a fictive data set from one player illustrated in long format.

474

475 Table 2. Estimated effects for the candidate variables from the univariable conditional logistic  
476 regression models.

477

478

479

480

481

482

483

484

485

486

487

488

489

490

491

492

493

494

495

496

497

498

499

500

501

502

503

504

505

506

507

508

509

510